## A Typical Sherpa Session

(The boiled-down version.)

#### The user:

- reads in data (and sets filters, etc.);
- builds model expressions;
- chooses a statistic;
- fits the model expressions to the data, one at a time;
- compares the results of the fits in order to select a best-fit model; and
- estimates the errors for the best-fit model parameters.

## Choosing a Statistic

(So many choices, so little guidance.)

A key feature of Sherpa is its large array of statistics appropriate for analyzing Poisson-distributed (i.e. counts) data.

- Statistics based on  $\chi^2$ :
  - CHI GEHRELS
  - CHI DVAR
  - CHI MVAR
  - CHI PARENT
  - CHI PRIMINI
- $\bullet$  Statistics based on the Poisson likelihood  $\mathcal{L}$ :
  - -CASH
  - BAYES

If the data are not Poisson-distributed (e.g. fluxes), then alternatives include:

- least-squares fitting: setting all variances to one; or
- providing errors in an input file.

# $\chi^2$ -Based Statistics

The  $\chi^2$  statistic is

$$\chi^2 \equiv \sum_i \frac{(D_i - M_i)^2}{\sigma_i^2},$$

where

- $D_i$  represents the observed datum in bin i;
- $M_i$  represents the predicted model counts in bin i; and
- $\sigma_i^2$  represents the variance of the sampling distribution for  $D_i$ .

$$\chi^2$$
 Statistic  $\sigma_i^2$ 

GEHRELS 
$$[1 + \sqrt{D_i + 0.75}]^2$$

$${ t DVAR}$$

$$exttt{MVAR} \qquad \qquad M_i$$

PARENT 
$$\frac{\sum_{i=1}^{N} D_i}{N}$$

PRIMINI 
$$M_i$$
 from previous best-fit

# Likelihood-Based Statistics

The CASH statistic is

$$C \equiv 2\sum_{i} [M_i - D_i \log M_i] \propto -2 \log \mathcal{L},$$

where

- $D_i$  represents the observed datum in bin i;
- $M_i$  represents the predicted model counts in bin i; and
- $\mathcal{L} = \prod_i \frac{M_i^{D_i}}{D_i!} \exp(-M_i).$

### Statistics: Caveats

(Potholes on the road to publication.)

Things to remember when using  $\chi^2$ :

- $\chi^2$  is an approximation of  $\log \mathcal{L}$  in the Gaussian (high-counts) limit. So...
- All estimations of variance (except GEHRELS) assume a Gaussian sampling distribution, not Poisson. Hence the number of counts in each bin should be  $\geq 5$ .
- CHI GEHRELS works with low-count data, but does not generally follow the  $\chi^2$  distribution: best fits are often "too good."
- And  $\chi^2$  is a biased estimator.

Things to remember when using CASH or BAYES:

- In the limit of high counts,  $\Delta C \sim \Delta \chi^2$ .
- Likelihood estimators are unbiased. But...
- Background subtraction is *not* allowed.
- There is no "goodness-of-fit" measure.
- And negative model amplitudes are *not* allowed.

## A Demonstration of Bias

- Using the *Sherpa* utility **FAKEIT**, we simulated 500 datasets from a constant model with amplitude 100 counts.
- We then fit each dataset with a constant model, recording the inferred amplitude.

| Statistic   | Average Amplitude |
|-------------|-------------------|
| CHI GEHRELS | 99.05             |
| CHI DVAR    | 99.02             |
| CHI MVAR    | 100.47            |
| CHI PARENT  | 99.94             |
| CHI PRIMINI | 99.94             |
| CASH        | 99.98             |
|             |                   |

## Optimization in Sherpa

Optimization is the action of minimizing  $\chi^2$  or  $-\log \mathcal{L}$  by varying the thawed parameters of the model. The user may choose between several optimization methods in Sherpa, including ones which:

- Find the local minimum.
  - POWELL
  - SIMPLEX
  - LEVENBERG-MARQUARDT

These algorithms are not computationally expensive, but they are also not appropriate for finding the global minimum of a complex statistical surface when starting from a random point.

- Attempt to find the global minimum.
  - GRID and GRID-POWELL
  - MONTE and MONTE-POWELL
  - SIMULATED ANNEALING

These are computationally intensive algorithms which are useful for searching complex statistical surfaces, starting from a random point.

## Optimization: Powell

**POWELL** is *Sherpa*'s default optimizer.

- It is a direction-set method in which initially, the chosen statistic is minimized by varying each parameter in turn while holding all other parameter values fixed.
- Advantages:
  - no gradient calculation
  - robust
    - \* can find local minima even on complex surfaces
    - \* can be used with all statistics
- Disadvantage:
  - relatively slow

## **Optimization:** Simplex

- The vertices of a simplex are reflected and/or contracted until the local minimum is bracketed.
- Advantages:
  - no gradient calculation
  - can find local minima even on complex surfaces
  - faster than POWELL
- Disadvantage:
  - exhibits a tendency to converge before reaching minima

## Optimization: Levenberg-Marquardt

• Approach the minimum taking steps of size  $\delta \vec{\theta}$ , computed by solving the set of linear equations:

$$\sum_{j=1}^{n} \alpha_{i,j} (1 + \lambda_{i,j}) \delta \theta_j = \beta_i,$$

where

$$\alpha_{i,j} = \sum_{k=1}^{n} \frac{1}{\sigma_k^2} \left[ \frac{\partial M(\vec{\theta})}{\partial \theta_i} \frac{\partial M(\vec{\theta})}{\partial \theta_j} \right] ,$$

$$\beta_i = -\frac{1}{2} \frac{\partial \chi^2}{\partial \theta_i} ,$$

and  $\lambda_{i,j}$  is a numerical factor, non-zero when i = j.

- Advantage:
  - fast
- Disadvantages:
  - requires gradient calculation
  - less robust in complex parameter spaces
  - appropriate for use with  $\chi^2$  statistics only
- Enhancements made in CIAO 2.1:
  - works correctly during simultaneous fits of source and background data
  - works correctly with double-precision data

## Confidence Intervals and Regions

(What are the errors on my parameters?)

- In frequentist statistics, the data are the random variables. Thus to estimate confidence intervals, new datasets need to be repeatedly simulated, either from the best-fit model or from the data themselves.
- A distribution of parameter values is generated by fitting the model to each simulated dataset.
- The central 68% of the parameter values can be called the  $1\sigma$  confidence interval.
- Simulations are computationally expensive. If:
  - the  $\chi^2$  or  $\log \mathcal{L}$  surface in parameter space is approximately shaped like a multi-dimensional paraboloid, and
  - the best-fit point is sufficiently far from parameter space boundaries,

then we may achieve good estimates of confidence intervals by examining the  $\chi^2$  or  $\log \mathcal{L}$  surface itself.

# Confidence Intervals and Regions: Uncertainty

- Vary a parameter's value, while holding the values of all other parameters to their best-fit values, until the fit statistic increases by some preset amount from its minimum value (e.g.  $\Delta \chi^2 = 1$  for  $1\sigma$ ).
- Gives correct results if and only if:
  - the statistic surface is "well-behaved"
  - there are no correlations between parameters
- Advantage:
  - fast
- Disadvantage:
  - errors are generally underestimated
- The user can visualize fit statistics as a function of parameter value using INTERVAL-UNCERTAINTY.
- The user can visualize two-dimensional confidence regions using REGION-UNCERTAINTY.

# Confidence Intervals and Regions: Projection

- Vary a parameter's value, while allowing the values of all other parameters to float to new best-fit values, until the fit statistic increases by some preset amount from its minimum value (e.g.  $\Delta \chi^2 = 1$  for  $1\sigma$ ).
- Gives correct results if and only if:
  - the statistic surface is "well-behaved"
- Advantages:
  - more accurate than UNCERTAINTY
  - provides a relatively inexpensize means of surface visualization
- Disadvantages:
  - no more accurate than the faster COVARIANCE
- The user can visualize fit statistics as a function of parameter value using INTERVAL-PROJECTION.
- The user can visualize two-dimensional confidence regions using REGION-PROJECTION.

# Confidence Intervals and Regions: Covariance

•  $1\sigma$  confidence intervals are given by  $\sqrt{C_{i,i}}$ , where

$$C_{i,j} = I_{i,j}^{-1},$$

and I, the information matrix computed at the best-fit point, is

$$I_{i,j} \equiv \frac{1}{2} \frac{\partial^2 \chi^2}{\partial p_i \partial p_j} \text{ or } \frac{1}{2} \frac{\partial^2 C}{\partial p_i \partial p_j} \text{ or } \frac{\partial^2 B}{\partial p_i \partial p_j}.$$

- Gives correct results if and only if:
  - the statistic surface is "well-behaved"
- Advantage:
  - fast
- Disadvantages:
  - the only computations are near the best-fit point, so not useful for surface visualization
  - involves matrix inversion, which can fail

# Example with a Well-Behaved Parameter Space

sherpa> fit powll: v1.2

powll: initial function value = 8.22297E+01

powll: converged to minimum = 6.27050E+01 at iteration = 7

powll: final function value = 6.27050E+01

p.c0 56.2579 p.c1 0.11117 p.c2 -0.00119999

sherpa> uncertainty

Computed for uncertainty.sigma = 1

| Parameter Name | Best-Fit    | Lower Bound  | Upper Bound  |
|----------------|-------------|--------------|--------------|
| p.c0           | 56.2579     | -0.865564    | +0.864461    |
| p.c1           | 0.11117     | -0.0148228   | +0.0148038   |
| p.c2           | -0.00119999 | -0.000189496 | +0.000189222 |

sherpa> projection

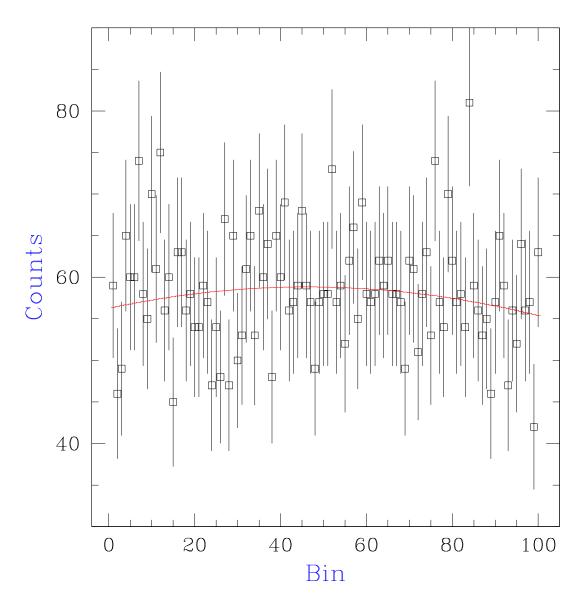
Computed for projection.sigma = 1

| Parameter Name | Best-Fit L  | ower Bound | Upper Bound |
|----------------|-------------|------------|-------------|
| p.c0           | 56.2579     | -0.120684  | +2.64497    |
| p.c1           | 0.11117     |            | +0.120703   |
| p.c2           | -0.00119999 |            | +0.00114976 |

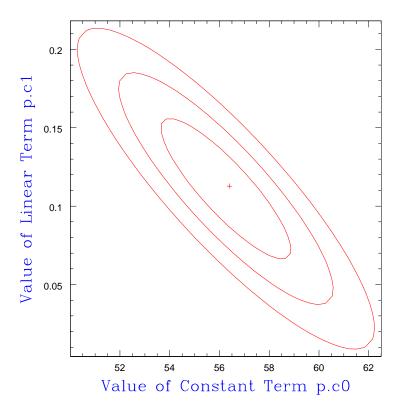
sherpa> covariance

Computed for covariance.sigma = 1

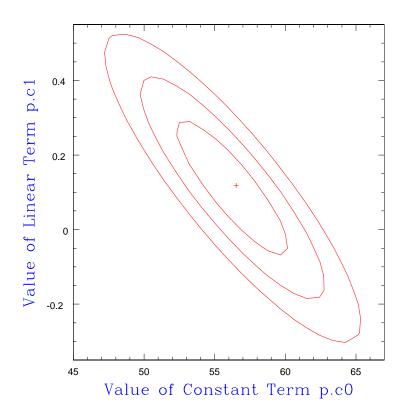
| Parameter Name | Best-Fit    | Best-Fit Lower Bound |             |
|----------------|-------------|----------------------|-------------|
| p.c0           | 56.2579     | -2.64786             | +2.64786    |
| p.c1           | 0.11117     | -0.121023            | +0.121023   |
| p.c2           | -0.00119999 | -0.00115675          | +0.00115675 |



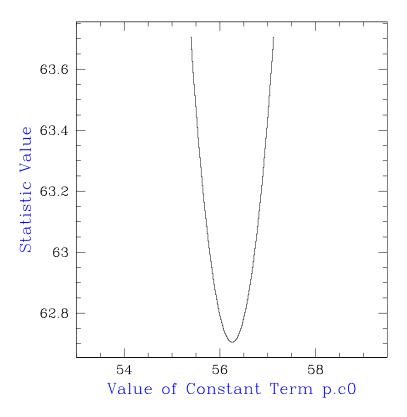
#### Confidence Region - Uncertainty



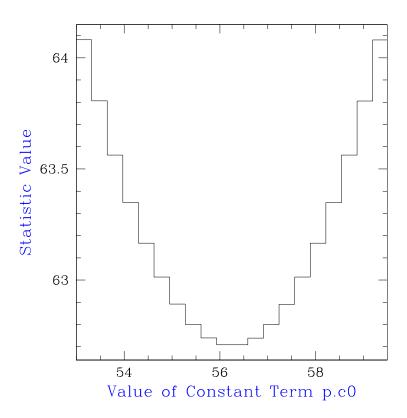
Confidence Region - Projection



#### Interval - Uncertainty



Interval - Projection



## Credible Intervals and Regions

(Bayesian methodology in the tiniest of nutshells.)

• In Bayesian methodology, credible intervals and regions are computed directly from the  $\chi^2$  or  $\log \mathcal{L}$  surface, using Bayes' theorem:

$$p(\vec{\theta}|D) \; = \; p(\vec{\theta}) \frac{p(D|\vec{\theta})}{p(D)} \, , \label{eq:posterior}$$

where

- $-p(D|\vec{\theta})$  is the likelihood of the data D given  $\vec{\theta}$ , the vector of model parameter values
  - (i.e.  $\log \mathcal{L}$  or  $\exp(-\chi^2/2)$ )
- $-p(\vec{\theta})$  is the prior for  $\vec{\theta}$
- $-p(\vec{\theta}|D)$  is the posterior for  $\vec{\theta}$
- -p(D) is an ignorable normalization constant
- The ability to specify priors is not yet included in *Sherpa*.

## Credible Intervals and Regions

• To estimate credible intervals, one marginalizes over nuisance parameters, e.g.:

$$p(\theta_1|D) = \int_{\theta_2} d\theta_2 \cdot \cdot \cdot \int_{\theta_n} d\theta_n \ p(\vec{\theta}|D).$$

- The central 68% of the distribution  $p(\theta_1|D)$  is the  $1\sigma$  credible interval.
- The computation of credible intervals and regions can be computationally intensive if there are many free parameters.
- However, approximate techniques such as adaptive integration are coded in freely available software, such as **BAYESPACK** (by Genz).

## Likelihood-Based Statistics

The **BAYES** statistic is the posterior distribution for the source model parameters  $\vec{\theta}_S$ , with the background amplitudes in each (energy) bin  $\theta_{B,i}$  marginalized out:

$$B \equiv -p(\vec{\theta}_S|D) = -\sum_i \int_{\theta_{B,i}} d\theta_{B,i} p(\vec{\theta}_S, \theta_{B,i}|D)$$

If  $\theta_{B,i}$  is *constant* as a function of spatial location and/or time, then an analytic expression (not reproduced here) replaces the summation of integrals.

NOTE:  $\theta_{B,i}$  are *implicit* parameters, not user-defined!

#### How is this statistic different from CASH?

- 1. CASH makes no assumptions about the behavior of the background as a function of spatial location and/or time.
- 2. CASH performs no implicit marginalization.

#### New Methods of Parameter Estimation

(Or, what might go into CIAO 4.0...)

Markov Chain Monte Carlo (MCMC) is a well-developed method that works as both an optimizer and a parameter estimator.

- A Markov Chain is an ordered sequence of random variables  $\Theta$ ; the probability of sampling variable  $\Theta_i$  depends only upon  $\Theta_{i-1}$ .
- The Monte Carlo aspect is how possible  $\Theta_i$  are chosen: randomly.

To use MCMC, a *Sherpa* user would:

- specify a rule for how possible  $\Theta_i$  are chosen (e.g. select new random values for a subset of the thawed parameters);
- specify a rule for whether  $\Theta_i$  is used, or disregarded (e.g. the Metropolis algorithm: given a randomly selected number r,  $0 \le r \le 1$ , keep  $\Theta_i$  if

$$r < \min \left[1, \frac{\mathcal{L}(\Theta_i)}{\mathcal{L}(\Theta_{i-1})}\right];$$

• and specify a stopping rule.

The central 68% of the selected parameter values define the  $1\sigma$  credible/confidence interval.

## Model Comparison Tests

(Which of my models is the best one?)

These do not yet exist in *Sherpa*. They compare directly compare two models,  $M_0$  and  $M_1$ , to yield either:

- The frequentist test significance,  $\alpha$ , that represents the probability of selecting the alternative (more complex) model  $M_1$  when in fact the null hypothesis  $M_0$  is correct; or
- The Bayesian odds, which is the ratio of model posterior probabilities for  $M_1$  and  $M_0$ :

$$O_{10} = \frac{p(M_1|D)}{p(M_0|D)}$$

In simple situations, the model posterior probability is determined by determining the integral of L over all parameter space.

## Model Comparison Tests

Standard model comparison tests include:

• The Maximum Likelihood Ratio (MLR) test:

$$\alpha_{\chi^2 \mathrm{MLR}} = \int_{\Delta \chi^2}^{\infty} d\chi^2 p(\Delta \chi^2 | \Delta N_{\theta}) ,$$

where  $\Delta N_{\theta}$  is the number of additional thawed model parameters in model  $M_1$ .

• The F-test:

$$\alpha_F = \int_F^{\infty} dF \ p(F|\Delta N_{\theta}, n - N_{\theta,1})$$

$$= I_{\frac{n-N_{\theta,1}}{n-N_{\theta,1} + (\Delta N_{\theta})F}} \left(\frac{n - N_{\theta,1}}{2}, \frac{\Delta N_{\theta}}{2}\right) ,$$

where n is the number of bins in the fit and  $N_{\theta,1}$  is the total number of thawed parameters in model  $M_1$ , I is the incomplete beta function, and F is the F-statistic

$$F = \frac{\Delta \chi^2}{\Delta N_{\theta}} / \frac{\chi_1^2}{(n - N_{\theta,1})}.$$

• Computation of the Bayesian odds using the Laplace approximation, valid for "well-behaved" surfaces. This approximation yields an analytic formula (not reproduced here) that allows the odds to be computed from  $\Delta \log \mathcal{L}$ ,  $\Delta N_{\theta}$ , the covariance matrices associated with both models, and the value of the priors at the best-fit points.

## Other Future Enhancements to Sherpa

- In convolution and optimization:
  - Treating pile-up.
  - Adding a convolution operator.
  - Adding the ability to use responses directly input from Fits Embedded Function (FEF) files when fitting models.
- In two-dimensional image analysis:
  - Being able to simultaneously fit source and background regions without inputting the background as a separate dataset.
  - Adding the ability to use exposure maps.
  - Extending flux calculations to two dimensions.
- In higher-dimensional data analysis:
  - Improving multi-axis fitting with functionals.
  - Adding visualization of data projected to one or two dimensions.

#### • And:

- Enhancing the capabilities of **GUIDE** to make it easier both to fit a sequence of individual lines and to perform differential emission measure fits.